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# Using multiple methods to assess detection probabilities of riparian-zone anurans: implications for monitoring

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# Abstract

*Context.* Both manual call surveys (MCS) and visual encounter surveys (VES) are popular methods used to monitor anuran populations. Recent statistical developments, specifically the development of occupancy models that permit the use of data from various survey methods to assess method-specific detection probabilities, provide a rigorous framework for evaluating the effectiveness of field methods.

*Aim.* To compare species-specific detection probabilities generated by MCS and VES and to evaluate the effectiveness of these methods throughout the activity season of several riparian-zone anuran species.

*Methods.* During 2010 and 2011, we sampled 21 sites along the Broad and Pacolet Rivers, in South Carolina, USA, using MCS and VES. Anuran species were surveyed across three seasons (fall, spring and summer) each year.

*Key results.* For six species, MCS resulted in a higher mean probability of detection, whereas VES resulted in a higher mean probability of detection for four species. In addition, survey date was an important influence on detection probability of most anurans when using MCS, but largely unimportant when employing VES.

*Conclusions.* Our findings indicated that VES are as effective as MCS for detecting some species of anurans, and for others, VES represent a more effective method. Furthermore, when using VES outside the breeding window, some anurans can be reliably detected, and in some cases, detected more easily than by using MCS.

*Implications.* We suggest that VES is a complimentary technique to MCS and a potentially important tool for population monitoring of anurans. VES can provide more flexibility for anuran researchers, as robust estimates of detection and occupancy can be obtained outside a narrow breeding window.

Additional keywords: active search, amphibian, manual call survey, South Carolina, survey method, visual encounter survey.

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# Introduction

Manual call surveys (MCS) are a popular method used to detect anurans for ecological, behavioural and conservationrelated investigations (e.g. Blair 1961; Woolbright 1985; Knutson et al. 2004; Price et al. 2005; Dorcas et al. 2010). In particular, using MCS to obtain anuran distribution data provides a relatively simple and cost-effective way to monitor populations across time and space (Droege and Eagle 2005; Weir et al. 2005; Dorcas et al. 2010). A major assumption of MCS is that peak anuran breeding windows are well understood and surveys are well timed to coincide with anuran reproductive activities. However, a central concern in any survey method, including MCS, is that of imperfect detection (MacKenzie et al. 2002, 2006). Species detection probability is defined as the probability of detecting at least one individual of a focal species during a sampling occasion, given that individuals of the species are present in the area (MacKenzie et al. 2002).

detection does not necessarily confirm absence. Thus, estimates of site occupancy and abundance, and their relationships with measured covariates (e.g. habitat type) along with estimated probabilities of colonisation and local extinction will be biased unless methods are used that can account for imperfect detection (MacKenzie et al. 2009). During MCS, even within the peak breeding season for many species, calling does not occur each night and there are variations in calling behaviour because of abiotic and biotic conditions, which can lead to the incorrect inference of absence of a species (Gooch et al. 2006; Brander et al. 2007; Cook et al. 2011). In addition, recent studies have shown that false positive errors (i.e. a species that is absent but erroneously detected) during anuran call surveys positively bias occupancy estimates, even when false positive detections occur only 1% of the time (Miller et al. 2011, 2012). For these reasons, it is becoming increasingly common for MCS monitoring

Although detection of a species confirms its presence, lack of

programs to standardise survey methodologies and account for factors that influence detection probability (Pellet and Schmidt 2005; Weir *et al.* 2005; Dorcas *et al.* 2010).

Another common amphibian survey method is visualencounter surveys (or active searches; VES), where observers visually search for amphibians in a designated area for a prescribed amount of time (Crump and Scott 1994). VES may hold distinct advantages over MCS. For example, some species vocalise only in response to heavy rains (e.g. spadefoot toads, Scaphiopus spp.), call infrequently (e.g. gopher frogs, Lithobates capito), have relatively short breeding seasons (wood frogs, Lithobates sylvaticus), are uncommon (e.g. Pellet and Schmidt 2005), or may not be audible over the louder higher-pitch calls of their contemporaries (Doan 2003; Droege and Eagle 2005). As with MCS, VES are based on the assumption that all individuals are equally detectable; other major assumptions are that there are no observer-related biases and that individuals are recorded once during a survey (Vonesh et al. 2010). Studies have used VES to assess the habitat use by focal species (Burbrink et al. 1998; Adams et al. 2011) and to compare the effectiveness of techniques such as cover-objects versus quadrat methodologies (Doan 2003; Grover 2006) and diurnal versus nocturnal searches (Heard et al. 2006), but direct comparisons with MCS have not been conducted. Furthermore, knowledge of how time of year influences detection probabilities of anurans observed via VES is not as well understood as it is for MCS.

It has become increasingly common for amphibian researchers to use multiple methods to detect species (Brown et al. 2007; Mattfeldt and Grant 2007; Dahl et al. 2009; Farmer et al. 2009; Balas et al. 2012). Some studies have expanded MCS to include vocal imitation of advertisement calls, in an attempt to stimulate male frogs and improve detection along with other VES methods such as nocturnal spotlight surveys and diurnal visual searches (e.g. Heard et al. 2006). Other studies examining optimal survey design in the tropics have used a combination of pitfall and funnel traps, along with MCS and VES, and automated tape recording of anuran calls (e.g. Parris et al. 1999; Rödel and Ernst 2004). Survey method is one of the most important considerations when designing studies. For example, when little is known about the natural history of a species, incorporating several methods along with covariates of interest (e.g. day of year) can provide estimates of detection probability, which allows researchers to identify the most appropriate sampling method. If detection probability is influenced by the time of year or weather conditions, for example, survey protocols can be adjusted to minimise the chance of false absences. Traditionally, researchers using multiple survey methods either combined data from each method to generate detection probabilities (i.e. Price et al. 2011), or separately generated detection probabilities for each method to compare their utility (Bailey et al. 2004; Mattfeldt and Grant 2007). Using multiple, single-method analyses or combined methods analysis ignores information from all but the focal detection method (Nichols et al. 2008). Recent advances in modelling circumvent some of the drawbacks of either combining sampling data or conducting single-method comparisons. Specifically, multi-method models permit simultaneous use of data from all methods for inference about method-specific detection probabilities (Nichols et al. 2008).

In the present study, our objectives were to evaluate two common methods for surveying anurans, namely, MCS and VES, by using a statistical model that permits incorporation of both methods into a single model (Nichols *et al.* 2008). We examined the generated detection probabilities for anuran species by using both methods and then examined how day of year influenced the effectiveness of each method. Given the variation in life history of many anurans, for some species, we expected that VES would be comparable to MCS and sometimes preferable to MCS, and that during certain times of the year, one method would be preferable over the other.

## Materials and methods

# Study area

Study sites (n=21) were located along the Broad and Pacolet Rivers, in the Piedmont region of north-central South Carolina (Fig. 1). We used a geographic information system (ArcGIS 10.0; Environmental Systems Research Institute, Redlands, CA, USA), with layers from the National Wetland Inventory (http://www. fws.gov/wetlands/, verified 20 July 2013), and the 2006 National Land Cover Database (Fry et al. 2011) to identify ~200 riparian wetlands within our study area. We generated a circular buffer of 1-km radius around each site, a distance that encompasses a majority of the core terrestrial habitat used by most anuran species (Semlitsch and Bodie 2003), and chose final study sites on the basis of spatial independence (i.e. non-overlapping 1-km-radius circular buffers). After ground-truthing, we determined that 21 sites were sufficiently accessible for time-constrained active searches (Fig. 1; see Eskew et al. (2012) and Hunt et al. (2013) for more information on study site selection).

#### Data collection

During 2010 and 2011 we surveyed each site using MCS (Dorcas et al. 2010) and time-constrained VES (Crump and Scott 1994) to document anurans. During 2010, three anuran calling surveys were conducted in both spring and summer at every site and during 2011 three anuran calling surveys were conducted in winter, spring and summer at every site; these seasons corresponded to the peak breeding windows for species in our study (Table 1). In total, 15 MCS were conducted at each study site. Manual call surveys were conducted between 1845 hours and 0100 hours by two experienced anuran surveyors listening independently for 5 min and recording all species heard; any differences in species observed were reconciled before leaving the study site and questionable species identifications were eliminated (Eskew et al. 2012). In total, there were four experienced observers participating in the study; however, only two observers participated in each survey. During each MCS, the number of cars passing by was recorded, because this may interfere with the ability to detect frogs. The majority of sites were located in rural areas and had little to no road traffic and, in cases with moderate traffic, observers waited to begin the survey until all cars has passed. Some survey sites were located nearer to dams than others (see Eskew et al. 2012); however, the distance was not great enough to influence observer ability to detect calling frogs (i.e. the noise level was always relatively low), and we found that water releases were infrequent relative to the timing of our surveys.

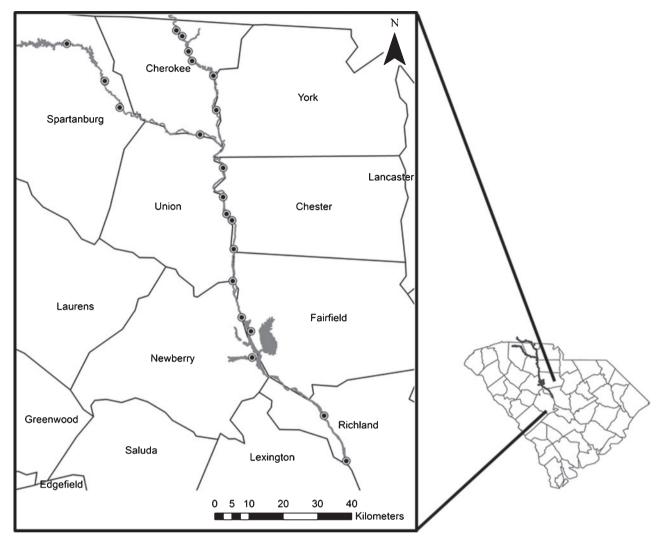


Fig. 1. Location map depicting South Carolina, USA. Inset shows locations of each anuran study site along both the Pacolet (upper left fork) and Broad River in South Carolina, USA; county boundaries are delineated on the South Carolina outline, and labelled on the study site inset.

Table 1. Proportion of sampling units where each species was observed (i.e. naïve estimate) and the number of times each
species was detected, for each survey method (i.e. visual encounter surveys (VES) and manual calling surveys (MCS)
In total, 12 repeat surveys were conducted for each survey method

Species	Common name	Naïve estimate (n	o. of detections)	Breeding window
		Active search	Call survey	
A. americanus	American toad	0.48 (16)	0.95 (40)	February to April
A. fowleri	Fowler's toad	0.90 (94)	1.00 (128)	April to July
A. terrestris	Southern toad	0.05 (1)	0.14 (4)	Late February to May
A. crepitans	Northern cricket frog	0.57 (86)	0.71 (66)	April to August
H. cinerea	Green treefrog	0.38 (20)	0.76 (65)	April to September
H. chrysoscelis	Cope's grey treefrog	0.90 (53)	0.95 (66)	April to August
P. feriarum	Upland chorus frog	0.43 (19)	0.90 (47)	November to early April
P. crucifer	Spring peeper	0.29 (11)	0.95 (102)	November to April
G. carolinensis	Eastern narrow-mouth toad	0.29 (18)	0.19 (4)	April to October
L. catesbeianus	Bullfrog	0.57 (33)	0.67 (52)	April to August
L. clamitans	Green frog	0.67 (39)	0.86 (45)	April to August
L. palustris	Pickerel frog	0.43 (13)	0.33 (12)	February to early April
L. sphenocephalus	Southern leopard frog	0.81 (79)	0.62 (32)	December to early April
S. holbrookii	Eastern spadefoot	0.10 (3)	0.00(0)	Year round

In addition, time-constrained VES were conducted at each site. During 2010, VES were conducted in summer (n=3) and fall (n=2) and, during 2011, VES were conducted in spring (n=3), summer (n=3) and fall (n=2) at each site; we assumed that we would detect focal species during these seasons, when activities such as breeding, ovipositing and foraging can be observed (Dorcas and Gibbons 2008). In total, 13 VES were conducted at each study site. Each survey consisted of two persons independently searching the site for 30 min to identify anuran species. A few searches were made by one person and these lasted 1 h. Each VES was conducted during daylight hours (i.e. 0900-1900 hours) and as much variation in microhabitat was sampled as possible, including surface cover objects (Vonesh et al. 2010). Search area was restricted to areas that we could hear calling amphibians from during MCS surveys. Anurans heard vocalising during VES were documented, along with any visual observations of species. All observed life-history stages were included during VES; however, the majority of encompassed adult and metamorphosing observations individuals (as compared with observations of tadpoles). For both survey methods, we recorded the number of days since rainfall and included this as a sampling covariate.

## Data analyses

We used the model Nichols et al. (2008) developed for multiple detection methods that allows simultaneous use of data from all methods for inference about method-specific detection probabilities. This modelling approach applies to any situation in which multiple detection methods are used in the same locations and allows parameters to be modelled as functions of site or sample-specific covariates. The general model comprised one detection parameter,  $p_{t}^{s}$ , which is the probability of detection at occasion t by method s, given the sample unit is occupied and the species is present at the immediate sample station site, and two occupancy parameters, namely,  $\psi$ , which is the probability a unit is occupied, and  $\theta_t$ , which is the probability a species is present at the immediate sample site at occasion t, given the sample unit is occupied. The two occupancy parameters,  $\psi$  and  $\theta_t$ , permit the modelling of occupancy at two different spatial scales; however, for the present study, we focussed on the influence of survey method on detection probability and held the two occupancy parameters constant.

For each species, we used data from 12 VES conducted in summer, fall and spring (i.e. peak activity period) paired with data from 12 MCS conducted during summer, winter and spring (i.e. peak breeding period). Because we conducted more call surveys than active searches, we eliminated call survey dates that fell outside a particular species' window of peak calling activity (Table 1). Peak anuran activity varies from species to species and we were able to fit each species into a general category of spring-summer-fall or fall-winter-spring activity. For example, the northern cricket frog (Acris crepitans) was placed in the spring/summer/fall activity window because it breeds from April to August (Table 1) and is active into the fall months. This exercise identified which three MCS surveys could be eliminated before data analysis, for each species, such that each survey method would have the same amount of effort (i.e. 12 VES and 12 MCS).

To obtain estimates of the probability of detection for each species for each method, we used the Nichols *et al.* (2008) model within program PRESENCE 5.7 (Hines 2006), which estimates parameters using maximum likelihood estimation methods. We evaluated the following eight models examining different influences on detection probability of anurans: (1) constant (null model with no covariates on detection probability), (2) survey method, (3) day of year ('date'), (4) number of days since rain ('rain'), (5) date and rain, (6) survey method and date, (7) survey method and rain and (8) survey method plus date and rain.

To obtain the most accurate parameter estimation, all covariates were standardised by calculating z-scores (i.e. the mean was subtracted from each value and then divided by the s.d.) before analysis. Model selection was based on Akaike information criteria (AIC; Burnham and Anderson 2002). We used AIC values adjusted for small samples sizes (i.e. AICc), and assessed fit for each model set by using the MacKenzie-Bailey goodness-of-fit test (MacKenzie and Bailey 2004). We conducted the test for 1000 bootstrap iterations on the most parameterised models in each model set to generate estimates of the overdispersion factor, ĉ, and used the ĉ value to ensure a conservative estimation of goodness-of-fit (i.e. if  $\hat{c} > 1$ , we used OAICc values adjusted for overdispersion; Burnham and Anderson 2002). We examined Akaike weights to determine the strength of evidence for each model (Burnham and Anderson 2002). We assumed models with higher weights and lower AIC values were better able to explain variation in data and selected the models with substantial empirical support, which included models within  $\Delta 2$  AICc (Burnham and Anderson 2002). In cases where there was no clear 'best' model, we computed model-averaged estimates (Buckland et al. 1997) for parameters of interest. We define 'best' to be a model within  $\Delta 2$  AICc which has >50% of the weight and where the remaining 50% of the weight is spread relatively thinly among the remaining candidate models that are also within  $\Delta 2$  AICc. Species-specific detection probabilities were derived using the inverse logittransformation (i.e.  $(\exp (\alpha)/(1 + \exp \alpha))$  of parameter estimates. We estimated lower and upper 95% confidence intervals using the delta method (Ver Hoef 2012). Because it was impractical to summarise AIC results for eight candidate model sets for each species, we present the top models for each species.

# Results

We detected the following 14 anuran species: American toad (Anaxyrus americanus), Fowler's toad (Anaxyrus fowleri), southern toad (Anaxyrus terrestris), northern cricket frog (A. crepitans), green treefrog (Hyla cinerea), Cope's grey treefrog (Hyla chrysoscelis), upland chorus frog (Pseudacris feriarum), spring peeper (Pseudacris crucifer), eastern narrowmouthed toad (Gastrophryne carolinensis), American bullfrog (Lithobates catesbeianus), green frog (Lithobates clamitans), pickerel frog (Lithobates palustris), southern leopard frog sphenocephalus) and (Lithobates eastern spadefoot (Scaphiopus holbrookii). The naïve occupancy and detection estimate varied among species for both MCS and VES (Table 1). Because S. holbrookii detections were so limited,

	Model	AICc	ΔQAICc	м	К	-2 log-likelihood			Parameter	Parameter estimate (s.e.)	ate (s.e.)	s.e.) Dav of vear	Munhar of days	of days
							Ψ-intercept	p-intercept	p-VES	p-MCS	p-VES	p-MCS	since rain p-VES p-M	p-MCS
A. americanus	$\Psi(.)^A$ , $\theta(.)^A$ , $p^B(Date)$	323.64	0.00	0.57	4	313.14	4.94	-2.87	I	I	-0.13	-1.38	I	Т
							(4.02)	(0.23)			(0.09)	(0.22)		
	Ψ(.), θ (.), p(Date, Rain)	325.54	1.90	0.22	9	307.54	4.87	-2.97	I	I	-0.09	-1.47	-0.64	0.19
	$\Psi(.), \theta(.), p(Method, Date)$	325.99	2.35	0.18	5	311.99	(4.00) 3.92	(0.27) -	-2.70	-3.23	(0.0) -0.07	(0.24) -1.64	(0.31) -	(0.17) -
							(2.54)		(0.26)	(0.42)	(0.08)	(0.34)		
	Model averaged						4.59 (3.62)	-2.29 (0.19)	-0.49 (0.05)	-0.58 (0.08)	-0.11 (0.09)	-1.41 (0.24)	-0.14 (0.07)	0.04 (0.04)
A. fowleri	Ψ(.), θ (.), p(Date, Rain)	648.63	0.00	0.73	9	630.63	2.12	-0.26	I	I	-0.51	1.24	0.06	0.84
	W() A() MMethod Date Rain)	650.58	1 05	200	٢	90 209	(0.50) 2.06	(0.15)	02.0-	<u>, 000</u>	(0.17)	(0.28) 1 10	(0.04) 0.04	(0.19) 0.86
	1 (·), V (·), P(************************************	0		1.0	-	0/	(0.48)		(0.17)	(0.22)	(0.17)	(0.30)	(0.15)	(0.20)
A. terrestris	Ψ(.), θ (.), p(.)	57.47	0.00	0.44	2	52.80	-0.95	-1.87	I	I	I	I	I	T
	Ψ(.), θ (.), n(Method)	57.98	0.51	0.34	ŝ	50.57	(0.7)	(1.19) _	-2.71	-1.30	I	I	I	I
					I	- - - -	(0.72)		(1.44)	(1.28)				
	Model averaged						-0.78	-0.82	-0.92	-0.44				
							(0.57)	(0.52)	(0.49)	(0.43)				
A. crepitans	$\Psi(.), \theta(.), p(Method, Date)$	530.38	0.00	0.55	ŝ	516.38	1.21	I	0.39	-0.49	0.10	0.73	I	I
	ΨΟ θ Ο n(Method)	531 49	111	0 31	"	524.08	(0.29) 1.08	I	(0.23)	(0.24)	(0.08)	(0.27)	I	I
		1.100		10.0	C	00.1170	(0.28)		(0.24)	(0.11)				
	Model averaged						1.00 (0.25)	I	0.34 (0.20)	-0.32 (0.17)	0.06 (0.04)	0.40 (0.15)	I	I
H. cinerea	$\Psi(.), \theta(.), p(Method, Date)$	354.84	0.00	0.54	2	340.84	0.70	I	-1.81	-0.98	-0.51	2.92	I	Ι
	Ψ(.), θ (.), p(Date)	355.21	0.37	0.45	4	344.71	(0.41) 0.87	-1.57	(0.33) -	(0.45) –	(0.32) -0.37	(1.13) 2.96	I	Ι
							(0.39)	(0.26)			(0.27)	(0.70)		
	Model averaged						0.77 (0.40)	-0.71 (0.12)	-0.98 (0.18)	-0.53 (0.24)	-0.44 (0.29)	2.90 (0.93)	I	I
H. chrysoscelis	Ψ(.), θ (.), p(Date)	526.82	0.00	0.58	4	516.32	2.02	-1.25	Ι	Ι	-0.92	0.81	Ι	Ι
	$\Psi(.), \theta(.), p(Method, Date)$	529.00	2.18	0.19	2	511.00	(1.00) 2.00 (1.08)	-	-1.41 (0.25)	-1.09 (0.25)	(0.27) -1.04 (0.32)	(0.27) 0.69 (0.27)		

Table 2. Best-supported models for detection of anuran species, based on QAIC<sub>c</sub> selection criteria The same eight candidate models were constructed for each species, from uncorrelated covariates. Best-supported models are in bold. AIC, Akaike information criteria;  $\Delta QAICc$ , difference in QAICc relative

(continued)	
Table 2.	

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Species	Model	AILC	AUAICC	≷	2	-7 log-likelihood			Paran Met	Parameter estimate (s.e.) Method Day	nate (s.e.) Day c	s.e.) Day of year	Number	Number of days
							W-intercept	p-intercept	p-VES	p-MCS	p-VES	p-MCS	since p-VES	since rain ES p-MCS
	Ψ(.), θ (.), p(Date, Rain)	529.02	2.20	0.19	9	515.02	2.05 (1.04)	-1.28 (0.20)	I	I	-0.93 (0.28)	0.82 (0.26)	-0.25 (0.17)	0.29 (0.16)
	Model averaged						1.94 (1.02)	-0.97 (0.16)	-0.27 (0.05)	-0.21 (0.05)	-0.91 (0.27)	0.76 (0.25)	-0.05 (0.04)	0.06 (0.03)
P. feriarum	$\Psi(.), \theta(.), p(Method, Date)$	306.14	0.00	0.97	ŝ	292.14	1.02 (0.32)	I	-2.51 (0.24)	-2.01 (0.66)	0.35 (0.24)	-3.33 (0.50)	I	I
P. crucifer	Ψ(.), θ (.), p(Method, Date)	412.65	0.00	0.85	S	398.65	0.87 (0.33)	I	-2.60 (0.34)	-1.20 (0.36)	0.46 (0.32)	-2.73	I	I
	Ψ(.), θ (.), p(Date)	416.41	3.76	0.13	4	405.91	0.71 (0.20)	-2.11 (0.24)		(ana)	(0.28) (0.28)	-4.95 (1.05)	I	I
G. carolinensis	$\Psi(.), heta$ (.), $p(Method)$	151.68	0.00	0.93	3	137.68	-0.41 (0 30)	I	-0.56	-2.37	I	I	I	I
L. catesbeianus	Ψ(.), θ (.), p(Date)	439.26	0.00	0.44	4	428.76	1.53	-1.28	-	-	-0.19	0.74	I	I
	$\Psi(.), \theta(.), p(Method)$	440.10	0.84	0.29	ŝ	432.69	(1.58) 1.58 (0.57)	(61.U) -	-1.43	-0.82	(c1.0) _	- -	I	I
	$\Psi(.), \theta(.), p(Method, Date)$	441.18	1.92	0.17	5	427.18	(0.52) (0.52)	I	(0.23) (0.23)	(0.25)	-0.15 (0.21)	0.58 (0.27)	I	I
	Model averaged						1.39 (0.47)	-0.56 (0.08)	-0.66 (0.11)	-0.42 (0.10)	-0.11 (0.10)	0.42 (0.16)		
L. clamitans	Ψ(.), θ (.), p(Date) Ψ(.), θ (.), p(Date, Rain)	441.53 443.76	0.00 2.23	0.58 0.19	4 0	431.03 425.76	1.78 (0.61) 1.79	-1.44 (0.19) -1.47	I I	I I	-0.29 (0.21) -0.31	0.72 (0.24) 0.74	- 0.42	- 0.03
L. palustris	Ψ(.), p(.)	200.71	0.00	0.65	7	196.04	(0.60) 1.01	(0.19) -2.26	I	I	(0.21)	(0.24)	(0.18)	(0.18)
·	Ψ(.), p(Method)	203.41	2.70	0.17	б	196.00	(0.67) 1.01 (0.70)	(0.41)	-2.22 (0.46)	-2.31 (0.46)	I	I	I	I
L. sphenocephalus	Ψ(.), θ (.), p(Method, Date)	484.14	0.00	0.78	S	470.14	1.63 (0.56)	I	-0.25	-2.49 (0.41)	-0.08	-1.10	I	I
	$\Psi(.), \theta(.), p(Method)$	488.04	3.90	0.11	б	480.63	(0.56)	I	-0.26 (0.22)	(0.23)	(00.0)	(0)	I	I
<sup>A</sup> Constant probabili <sup>B</sup> Probability of dete	<sup>A</sup> Constant probability of occupancy (i.e. $\Psi(.)$ , $\theta(.)$ ). <sup>B</sup> Probability of detection, which varies by model for each species, with or without inclusion of covariates that improve detection estimates.	ch species, v	vith or withc	out inclu	sion o	f covariates that imp	prove detection	estimates.						

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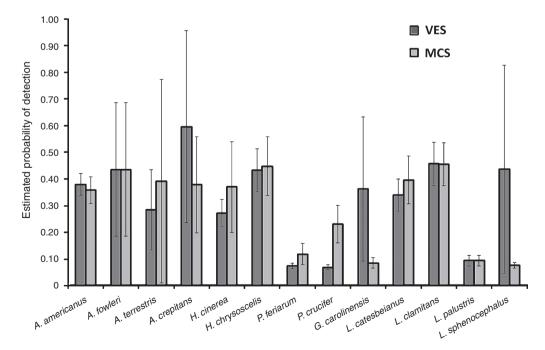
method-specific detection probabilities for this species were not possible to estimate.

The best supported model for 10 anuran species indicated that detection varied by survey method (i.e. VES or MCS; Table 2, Appendix 1); for four of these species (*P. feriarum*, *P. crucifer*, *G. carolinensis* and *L. sphenocephalus*), detection varied by survey method and, for six species, the best-supported model was determined from model averaging (*A. americanus*, *A. terrestris*, *A. crepitans*, *H. cinerea*, *H. chrysoscelis* and *L. catesbeianus*). For the remaining three species (*A. fowleri*, *L. palustris* and *L. clamitans*), survey method was unimportant and the best-supported model suggested either a constant probability of detection (*L. palustris*) or suggested that detection varied by date (*L. clamitans*) or date and days since rainfall (*A. fowleri*; Table 2).

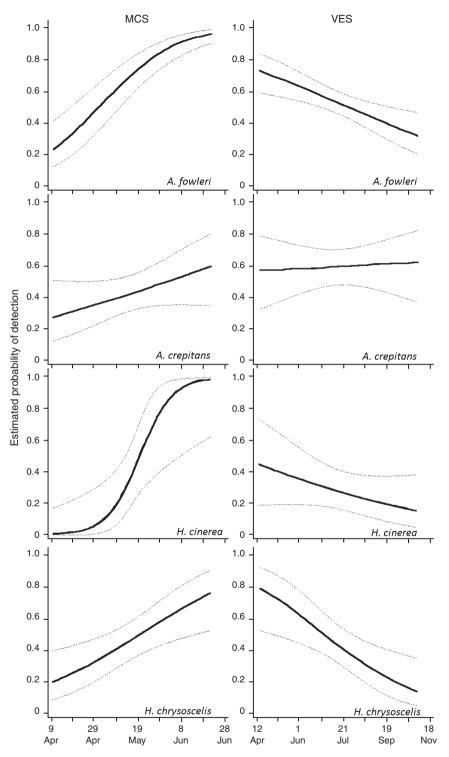
For six species, MCS resulted in a higher mean (s.e.) probability of detection than did VES (A. terrestris: MCS 0.39 (0.38) vs VES 0.28 0.15); H. cinerea: MCS 0.37 (0.17) vs VES 0.27 (0.05); H. chrvsoscelis: MCS 0.45 (0.11) vs VES 0.43 (0.08); P. feriarum: MCS 0.12 (0.04) vs VES 0.08 (0.01); P. crucifer: MCS 0.23 (0.07) vs VES 0.07 (0.01); and L. catesbeianus: MCS 0.40 (0.09) vs VES 0.34 (0.06); Fig. 2). Conversely, for four species, VES resulted in a higher mean probability of detection than did MCS (A. americanus: VES 0.38 (0.04) vs MCS 0.36 (0.05); A. crepitans: VES 0.60 (0.36) vs MCS 0.38 (0.18); G. carolinensis: VES 0.36 (0.27) vs MCS 0.09 (0.02); and L. sphenocephalus: VES 0.44 (0.39) vs MCS 0.08 (0.01); Fig. 2). For three species, however, both methods were equally effective and detection probabilities were identical (A. fowleri: 0.44 (0.25), L. clamitans: 0.46 (0.08) and L. palustris: 0.09 (0.02); Fig. 2).

In addition to survey method, date was an important factor influencing detection probability and was included in the best-supported model for 10 species (A. americanus, A. fowleri, A. crepitans, H. cinerea, H. chrysoscelis, P. feriarum, P. crucifer, L. catesbeianus, L. clamitans and L. sphenocephalus; Table 2). For all summer-breeding frogs except A. crepitans, detection probability during the VES decreased with increasing day of year (i.e. decreased from spring to fall), which corresponded to between 12 April and 18 November each year (Figs 3, 4). Specifically, from mid-April to mid-November of each year, the estimated probability of detection during VES increased from 0.58 to 0.60 for A. crepitans, and decreased from 0.40 to 0.30 for L. catesbeianus, from 0.78 to 0.32 for A. fowleri, from 0.45 to 0.19 for H. cinerea, from 0.73 to 0.48 for L. clamitans, and from 0.80 to 0.17 for H. chrysoscelis (Figs 3, 4). For all summerbreeding frogs, detection probability during the MCS calling window increased with increasing day of year, which corresponds to between 9 April and 28 June each year (Figs 3, 4). Specifically, from early April to late June of each year, the estimated probability of detection during MCS increased from 0.25 to 0.60 for A. crepitans, from 0.22 to 0.60 for L. catesbeianus, from 0.21 to 0.95 for A. fowleri, from 0.002 to 0.99 for *H. cinerea*, from 0.22 to 0.80 for *L. clamitans*, and from 0.20 to 0.78 for *H. chrysoscelis* (Figs 3, 4).

For all winter-breeding frogs, detection probability during the active-search survey window varied slightly with increasing day of year (Fig. 5); the estimated probability of detection during VES decreased from 0.50 to 0.40 for *L. sphenocephalus* and from 0.41 to 0.31 for *A. americanus*, and increased from 0.05 to 0.18 for *P. feriarum* and from 0.03 to



**Fig. 2.** Estimated conditional probability of detection  $(\pm 1 \text{ s.e.})$  for each anuran species along the Broad and Pacolet Rivers, South Carolina, USA. Estimates are based on repeat surveys using two different survey methods (MCS or VES).



**Fig. 3.** Effect of date by survey method (MCS or VES) on conditional detection probability of summer-breeding anuran species along the Broad and Pacolet Rivers, South Carolina, USA. Solid lines represent the relationship between date and detection probability and dashed lines are the 95% CIs for estimates of the covariate effect.

0.18 for *P. crucifer* (Fig. 5). For all winter-breeding frogs, detection probability during the calling-survey time span decreased with increasing day of year (Fig. 5). Specifically,

from mid-February to mid-May of each year, the estimated probability of detection during MCS decreased from 0.38 to 0.02 for *L. sphenocephalus*, from 0.98 to 0.01 for *P. feriarum*,

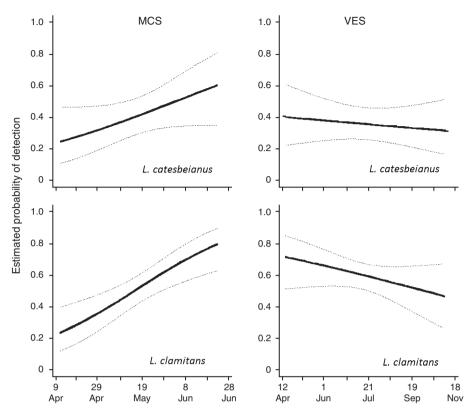


Fig. 4. Effect of date by survey method (MCS or VES) on conditional detection probability of summerbreeding anuran species along the Broad and Pacolet Rivers, South Carolina, USA. Solid lines represent the relationship between date and detection probability and dashed lines are the 95% CIs for estimates of the covariate effect.

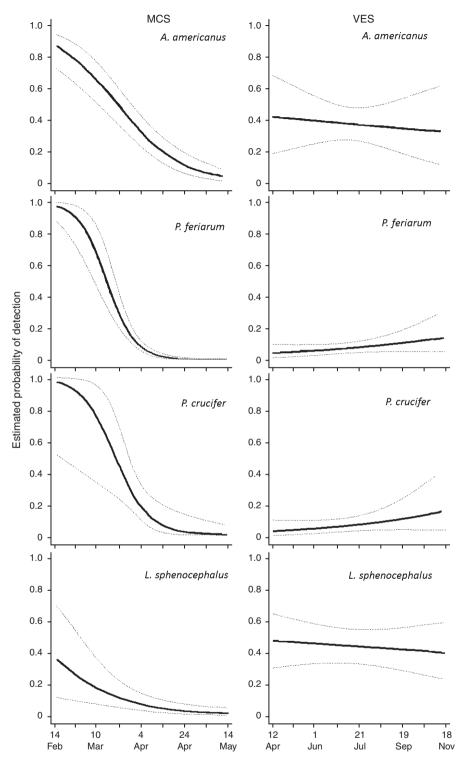
from 0.89 to 0.05 for *A. americanus*, and from 0.98 to 0.01 for *P. crucifer* (Fig. 5).

### Discussion

Our study has provided insight for monitoring anurans by implementing a modelling framework where data from both MCS and VES are evaluated simultaneously, resulting in method-specific detection probabilities. Our results provided strong evidence of variation in anuran detection probabilities among two different detection methods. Specifically, for 10 of 13 species, the best-supported model indicated that detection varied by survey method. Consistent with previous studies (e.g. Gooch et al. 2006; Steelman and Dorcas 2010), calling surveys were effective at detecting a wide range of anurans. Visual encounter surveys, however, were also effective at detecting a wide range of anuran species, and several species had a higher detection probability with VES than with MCS. Visual encounter surveys are likely to be under-used as a monitoring tool for anurans, perhaps because of a lack of information on their utility. Instead, VES are often used along with other methods in surveys for focal or rare species where the intent is geared toward complete coverage of all life stages to increase detection probability (e.g. Drost and Fellers 1996; Burbrink et al. 1998; Lips 1999; Lehtinen and Galatowitsch 2001).

Although both MCS and VES are effective methods, for certain species in our study, one method was most effective.

Specifically, VES represent a more effective method for detecting A. crepitans, G. carolinensis and L. sphenocephalus, which is perhaps reflective of variation in species life history. For example, G. carolinensis and L. sphenocephalus, like most anurans, can often be observed via vocalisations made during breeding; however, the breeding windows are wide (particularly for L. sphenocephalus), calling by both species is sporadic, and for G. carolinensis, calling is usually initiated only by heavy rains (Brandt 1936; Dodd 2013). In addition, calls of both G. carolinensis and L. sphenocephalus are easily masked by louder species (e.g. H. cinerea and P. crucifer; Conant and Collins 1998). Likely because both G. carolinensis and L. sphenocephalus are active during the non-breeding period in a wide variety of aquatic and terrestrial habitats (Mitchell and Lannoo 2005; Dodd 2013), VES was a more reliable method of detection. Acris crepitans, another species with higher detection probabilities during VES, is active and conspicuous during the non-breeding season. Several other anuran species might also be more reliably detected using VES if their vocalisations are difficult to observe and their non-breeding habits permit detection (i.e. little grass frogs (Pseudacris ocularis) or species that exhibit explosive and/or short-lived breeding periods (e.g. Scaphiopus spp., L. sylvaticus, L. capito). In these cases, precisely timing MCS during this brief window is challenging (Dostine et al. 2013) and often not ideal because anuran call surveys are less effective during heavy rain (Weir et al. 2005).



**Fig. 5.** Effect of date by survey method (MCS or VES) on conditional detection probability of winterbreeding anuran species along the Broad and Pacolet Rivers, South Carolina, USA. Solid lines represent the relationship between date and detection probability and dashed lines are the 95% CIs for estimates of the covariate effect.

We found that MCS represented a more effective detection method for *H. cinerea*, *P. feriarum*, *P. crucifer* and *L. catesbeianus*. These species are common, vocal, and calls can

be heard from a considerable distance (Dorcas *et al.* 2007). For a few species, however, both survey methods were equally effective (i.e. for *A. fowleri*, *L. clamitans* and *L. palustris*). In addition, given the marginally larger detection probabilities of A. americanus using VES and of H. chrysoscelis using MCS, for these species, we also consider both survey methods to be similarly effective. For both A. americanus and A. fowleri, MCS and VES detection probabilities were very similar: MCS are effective because males of each species produce loud easily recognisable calls (Dorcas et al. 2007) but VES are also effective, perhaps because of their conspicuous tadpoles which school in large groups (Lefcort 1998) and the tendency for large numbers of metamorphs to forage for several weeks around the edge of the aquatic habitat from which they emerged (Beck and Congdon 1999; Dodd 2013). Likewise, VES had detection probabilities similar to those of MCS for L. clamitans, L. palustris and H. chrysoscelis; for L. clamitans, this may be because when startled, they emit conspicuous alarm calls as they leap into the water (Dorcas et al. 2007) and L. palustris may call in any month of the year in the southern United States (Conant and Collins 1998). Finally, for H. chrysoscelis, its tendency to call sporadically during the daytime likely accounts for the similarity in effectiveness of both survey methods.

Date was an important factor influencing detection probability of anurans, particularly regarding call surveys. For summerbreeding anurans, detection probability steadily increased toward the end of the summer survey windows, and likewise, for winter-breeding anurans, detection probability was highest toward the beginning of the winter survey windows. Previous studies have documented strong patterns in seasonal breeding activity and breeding windows are well established (Droege and Eagle 2005). However, for VES, the potential influence of date to increase or decrease the effectiveness of this survey method is not well known. Our results indicated that for VES, for most species, detection probability was relatively constant throughout the sampling window (i.e. April-November each year), although for a few species (A. fowleri, H. chrysoscelis, H. cinerea and L. clamitans), detection probability steadily decreased into the fall months. Furthermore, although we showed that VES are better for some species, we also highlighted that VES can provide adequate occupancy estimation, even when conducted outside the breeding window. For example, for L. clamitans and A. fowleri, VES conducted in early April resulted in a 60% probability of detection and, during MCS, it was not until late June, toward the end of the breeding season, that detection probabilities reached 60%. Similarly, for H. chrysoscelis, VES conducted in early April resulted in an 80% probability of detection and for A. crepitans, detection probabilities remained ~60% during the VES survey window; for both of these species, detection probabilities comparable to 80% and 60%, respectively, did not occur until late June. Therefore, we suggest that spring and fall VES can detect species that vocalise mostly during the summer months.

#### Management implications

The goals of a particular study should dictate the appropriate survey method; however, results of the present study highlighted the importance of estimating method-specific detection probabilities, which should prove especially useful in the design of multispecies surveys. We recommend the use MCS for anurans with broad calling windows and conspicuous vocalisations, and VES for anurans that either vocalise sporadically or discreetly. Furthermore, we recommend the use of VES throughout the activity season, not just the breeding season, because we found that survey date was a largely unimportant predictor of detection using VES. Our results also suggested that VES can provide more flexibility for anuran researchers, because robust estimates of detection and occupancy can be obtained outside a narrow breeding window.

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## Appendix 1. Complete candidate model results for detection of anuran species along the Broad and Pacolet Rivers, South Carolina, USA The same eight candidate models were constructed for each species, from uncorrelated covariates. Best supported models are in bold. AIC, Akaike information criteria; ΔQAICc, difference in QAICc relative to the top model; w, model weight; ĉ, overdispersion parameter estimated by dividing model deviance of the observed model by mean bootstrap model deviance; K, number of parameters in the model. Day of year, day of year each survey was conducted; Days since rain, number of days since rain for a given survey; Method, anuran survey method (VES or MCS)

Species	Model	AICc	ΔQAICc	W	K	-2 log-likelihood	ĉ
A. americanus	$\Psi(.)^{A}, \theta(.)^{A}, p(Date)^{B}$	323.64	0.00	0.57	4	313.14	
	$\Psi(.), \theta(.), p(Date, Rain)$	325.54	1.90	0.22	6	307.54	
	$\Psi(.), \theta(.), p(Method, Date)$	325.99	2.35	0.18	5	311.99	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	329.43	5.79	0.03	7	306.81	1.39
	$\Psi(.), \theta(.), p(Method)$	347.12	23.48	0.00	3	339.71	
	$\Psi(.), \theta(.), p(Method, Rain)$	349.05	25.41	0.00	5	335.05	
	Ψ(.),θ(.),p(.)	356.29	32.65	0.00	2	351.62	
	Ψ(.),θ(.),p(Rain)	358.85	35.21	0.00	4	348.35	
A. fowleri	$\Psi(.), \theta(.), p(Date, Rain)$	648.63	0.00	0.73	6	630.63	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	650.58	1.95	0.27	7	627.96	0.63
	$\Psi(.), \theta(.), p(Date)$	667.23	18.60	0.00	4	656.73	
	$\Psi(.), \theta(.), p(Method, Date)$	667.80	19.17	0.00	5	653.80	
	$\Psi(.), \theta(.), p(Method, Rain)$	671.85	23.22	0.00	5	657.85	
	$\Psi(.),\theta(.),p(Rain)$	677.55	28.92	0.00	4	667.05	
	$\Psi(.), \theta(.), p(Method)$	689.56	40.93	0.00	3	682.15	
	Ψ(.),θ(.),p(.)	696.20	47.57	0.00	2	691.53	
4. terrestris	Ψ(.),θ(.),p(.)	57.47	0.00	0.44	2	52.80	
	$\Psi(.), \theta(.), p(Method)$	57.98	0.51	0.34	3	50.57	
	$\Psi(.), \theta(.), p(Method, Rain)$	61.25	3.78	0.07	5	47.25	
	$\Psi(.), \theta(.), p(Date, Rain)$	61.49	4.02	0.06	6	43.49	
	$\Psi(.), \theta(.), p(Rain)$	62.50	5.03	0.04	4	52.00	
	$\Psi(.), \theta(.), p(Date)$	62.93	5.46	0.03	4	52.43	
	$\Psi(.), \theta(.), p(Method, Date)$	64.13	6.66	0.02	5	50.13	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	65.60	8.13	0.01	7	42.98	1.10
4. crepitans	$\Psi(.), \theta(.), p(Method, Date)$	530.38	0.00	0.55	5	516.38	
	$\Psi(.), \theta(.), p(Method)$	531.49	1.11	0.31	3	524.08	
	Ψ(.),θ(.),p(.)	534.35	3.97	0.08	2	529.68	
	$\Psi(.), \theta(.), p(Method, Rain)$	536.55	6.17	0.03	5	522.55	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	537.34	6.96	0.02	7	514.72	2.78
	$\Psi(.), \theta(.), p(Date)$	537.95	7.57	0.01	4	527.45	
	$\Psi(.), \theta(.), p(Rain)$	538.50	8.12	0.01	4	528.00	
	$\Psi(.), \theta(.), p(Date, Rain)$	543.65	13.27	0.00	6	525.65	
H. cinerea	$\Psi(.), \theta(.), p(Method, Date)$	354.84	0.00	0.54	5	340.84	
	$\Psi(.), \theta(.), p(Date)$	355.21	0.37	0.45	4	344.71	
	$\Psi(.), \theta(.), p(Date, Rain)$	362.64	7.80	0.01	6	344.64	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	363.34	8.50	0.01	7	340.72	4.61
	$\Psi(.), \theta(.), p(Method)$	404.25	49.41	0.00	3	396.84	
	$\Psi(.), \theta(.), p(Method, Rain)$	410.30	55.46	0.00	5	396.30	
	Ψ(.),θ(.),p(.)	435.67	80.83	0.00	2	431.00	
	Ψ(.),θ(.),p(Rain)	441.22	86.38	0.00	4	430.72	
H. chrysoscelis	$\Psi(.), \theta(.), p(Date)$	526.82	0.00	0.58	4	516.32	
	$\Psi(.), \theta(.), p(Method, Date)$	529.00	2.18	0.19	5	511.00	
	$\Psi(.), \theta(.), p(Date, Rain)$	529.02	2.20	0.19	6	515.02	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	532.27	5.45	0.04	7	509.65	6.30
	Ψ(.),θ(.),p(.)	550.47	23.65	0.00	2	545.80	
	$\Psi(.),\theta(.),p(Rain)$	551.20	24.38	0.00	4	540.70	
	$\Psi(.), \theta(.), p(Method)$	551.30	24.48	0.00	3	543.89	
	$\Psi(.), \theta(.), p(Method, Rain)$	552.77	25.95	0.00	5	538.77	
P. feriarum	$\Psi(.), \theta(.), p(Method, Date)$	306.14	0.00	0.97	5	292.14	
	$\Psi(.), \theta(.), p(Method, Rain, Date)$	313.07	6.93	0.03	7	290.45	6.21
	Ψ(.),θ(.),p(Date)	319.39	13.25	0.00	4	308.89	
	$\Psi(.), \theta(.), p(Rain, Date)$	325.75	19.61	0.00	6	307.75	
	$\Psi(.), \theta(.), p(Method)$	384.65	78.51	0.00	3	377.24	

(continued next page)

Species	Model	AICc	ΔQAICc	W	К	-2 log-likelihood	ĉ
	$\Psi(.), \theta(.), p(Method, Rain)$	390.63	84.49	0.00	5	376.63	
	Ψ(.),θ(.),p(.)	395.97	89.83	0.00	2	391.30	
	$\Psi(.), \theta(.), p(Rain)$	401.13	94.99	0.00	4	390.63	
P. crucifer	$\Psi(.), \theta(.), p(Method, Date)$	412.65	0.00	0.85	5	398.65	
	$\Psi(.), \theta(.), p(Date)$	416.41	3.76	0.13	4	405.91	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	420.88	8.23	0.01	7	398.26	4.46
	$\Psi(.), \theta(.), p(Date, Rain)$	423.32	10.67	0.00	6	405.32	
	$\Psi(.), \theta(.), p(Method)$	431.35	18.70	0.00	3	423.94	
	$\Psi(.), \theta(.), p(Method, Rain)$	437.70	25.05	0.00	5	423.70	
	Ψ(.),θ(.),p(.)	537.28	124.63	0.00	2	532.61	
	Ψ(.),θ(.),p(Rain)	542.89	130.24	0.00	4	532.39	
G. carolinensis	$\Psi(.), \theta(.), p(Method)$	151.68	0.00	0.93	5	137.68	
	$\Psi(.), \theta(.), p(Method, Rain)$	158.10	6.42	0.06	7	135.48	
	$\Psi(.), \theta(.), p(.)$	160.65	8.97	0.01	3	153.24	
	$\Psi(.), \theta(.), p(Rain)$	166.02	14.34	0.00	5	152.02	
	$\Psi(.),\theta(.),p(Date)$	170.07	18.39	0.00	2	165.40	
	$\Psi(.), \theta(.), p(Rain, Date)$	175.78	24.10	0.00	4	165.28	
	$\Psi(.), \theta(.), p(Method, Date)$	175.80	24.12	0.00	4	165.30	
	$\Psi(.), \theta(.), p(Method, Rain, Date)$	183.19	31.51	0.00	6	165.19	2.06
l. catesbeianus	$\Psi(.), \theta(.), p(Date)$	439.26	0.00	0.44	4	428.76	
n careso crantas	$\Psi(.), \theta(.), p(Method)$	440.10	0.84	0.29	3	432.69	
	$\Psi(.), \theta(.), p(Method, Date)$	441.18	1.92	0.17	5	427.18	
	$\Psi(.), \theta(.), p(.)$	443.05	3.79	0.07	2	438.38	
	$\Psi(.), \theta(.), p(Method, Rain)$	445.78	6.52	0.02	5	431.78	
	$\Psi(.),\theta(.),p(Date, Rain)$	445.80	6.54	0.02	6	427.80	
	$\Psi(.), \theta(.), p(Rain)$	448.07	8.81	0.01	4	437.57	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	448.77	9.51	0.00	7	426.15	5.34
L. clamitans	$\Psi(.), \theta(.), p(Date)$	441.53	0.00	0.58	4	431.03	
	$\Psi(.), \theta(.), p(Date, Rain)$	443.76	2.23	0.19	6	425.76	
	$\Psi(.), \theta(.), p(Method, Date)$	444.74	3.21	0.12	5	430.74	
	$\Psi(.), \theta(.), p(.)$	446.71	5.18	0.04	2	442.04	
	$\Psi(.),\theta(.),p(Rain)$	447.78	6.25	0.03	4	437.28	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	448.21	6.68	0.02	7	425.59	4.98
	$\Psi(.), \theta(.), p(Method)$	448.89	7.36	0.01	3	441.48	
	$\Psi(.),\theta(.),p(Method, Rain)$	450.54	9.01	0.01	5	436.54	
L. palustris	$\Psi(.), \theta(.), p(.)$	200.71	0.00	0.65	2	196.04	
	$\Psi(.), \theta(.), p(Method)$	203.41	2.70	0.17	3	196.00	
	$\Psi(.), \theta(.), p(Date)$	205.15	4.44	0.07	4	194.65	
	$\Psi(.),\theta(.),p(Rain)$	205.32	4.61	0.06	4	194.82	
	$\Psi(.), \theta(.), p(Method, Date)$	206.76	6.05	0.03	5	192.76	
	$\Psi(.), \theta(.), p(Method, Rain)$	208.80	8.09	0.01	5	194.80	
	$\Psi(.), \theta(.), p(Date, Rain)$	211.37	10.66	0.00	6	193.37	
	$\Psi(.), \theta(.), p(Method, Date, Rain)$	214.34	13.63	0.00	7	191.72	3.22
L. sphenocephalus	$\Psi(.), \theta(.), p(Method, Date)$	484.14	0.00	0.78	5	470.14	
2. sphenocephulus	$\Psi(.), \theta(.), p(Method)$	488.04	3.90	0.11	3	480.63	
	$\Psi(.), \theta(.), p(Method, Rain, Date)$	488.62	4.48	0.08	7	466.00	6.61
	$\Psi(.), \theta(.), p(Method, Rain)$	491.11	6.97	0.02	5	477.11	0.01
	$\Psi(.), \theta(.), p(.)$	514.39	30.25	0.02	2	509.72	
	Ψ(.),θ(.),p(Date)	516.53	32.39	0.00	4	506.03	
	$\Psi(.), \theta(.), p(Rain)$	517.56	33.42	0.00	4	507.06	
	$\Psi(.), \theta(.), p(Rain, Date)$	521.38	37.24	0.00	6	503.38	

Appendix 1. (continued)

<sup>A</sup>Constant probability of occupancy (i.e.  $\Psi(.)$ ,  $\theta$  (.)). <sup>B</sup>Probability of detection, which varies by model for each species, with or without inclusion of covariates that improve detection estimates.